

TEDAS - Tail Event Driven ASset allocation

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S&P 500 Stocks

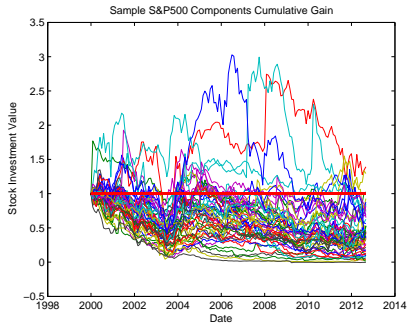


Figure 1: 50 random S&P 500 Sample Components' Cumulative Return: **94%** of stocks lost the value of the initial investment (**thick red line**)



Core & Satellites

Hedge funds, SDAX, MDAX and TecDAX constituents

- diversification - reduction of the portfolio risk
- construction - a more diverse universe of assets
- allocation - a higher risk-adjusted return.



Hedge Funds

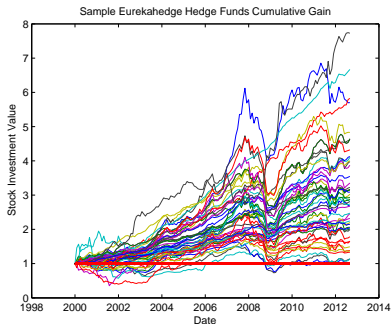


Figure 2: 50 EurekaHedge Hedge Funds Indices' Cumulative Return: **0%** of funds lost the value of the initial investment (**thick red line**)



Diversification

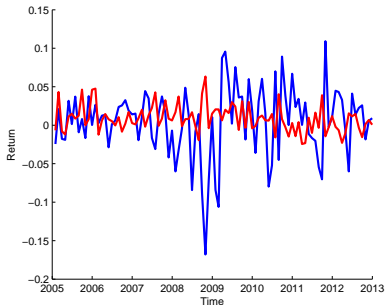


Figure 3: S&P 500 and Eureka Hedge North America Macro Hedge Fund Index monthly returns in 20050131-20121231



Traditional Assets/Hedge Funds

| Hedge Funds | US | UK | SW | GER | JAP |
|-----------------|-------|-------|-------|-------|-------|
| Conv. arb. | 0.10 | 0.08 | 0.07 | 0.10 | -0.02 |
| Dedic. sh. bias | -0.77 | -0.53 | -0.33 | -0.46 | -0.48 |
| Fix. inc. arb. | 0.10 | 0.13 | 0.01 | 0.08 | -0.10 |
| Glob. macro | 0.30 | 0.19 | 0.10 | 0.27 | -0.11 |
| Man. fut. | -0.10 | 0.02 | -0.09 | -0.03 | 0.03 |

Table 1: Correlation statistics for traditional asset class and hedge funds' indices; based on monthly data Jan. 1994 - Aug. 2001; table from Lhabitant (2002, p.164)

▶ TLND Hedge Funds' Strategies

▶ More



Tail Risk

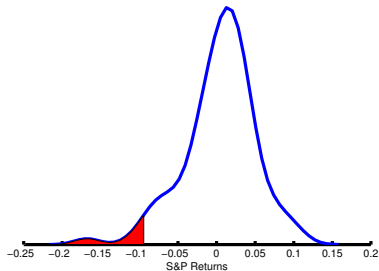


Figure 4: Estimated density of S&P 500 returns



The TLND challenge

- Tail dependence
- Large universe: $p > n$
- Non normality
- Dynamics



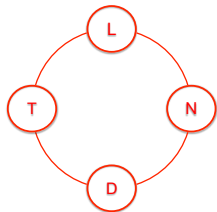
TEDAS Objectives

□ Hedge tail events

- ▶ Quantile regression
- ▶ Variable selection in high dimensions

□ Improve Asset Allocation

- ▶ Higher-order moments' optimization
- ▶ Modelling of moments' dynamics



Outline

1. Motivation ✓
2. TEDAS framework
3. Empirical Application
4. Conclusions
5. Technical Details

Tail Events

- $Y \in \mathbb{R}^n$ core log-returns; $X \in \mathbb{R}^{n \times p}$ satellites' log-returns, $p > n$
-

$$q_\tau(x) \stackrel{\text{def}}{=} F_{Y|X}^{-1}(\tau) = x^\top \beta(\tau) = \arg \min_{\beta \in \mathbb{R}^p} E_{Y|X=x} \rho_\tau\{Y - X^\top \beta\},$$

$$\rho_\tau(u) = u\{\tau - \mathbf{I}(u < 0)\}$$

- L_1 penalty $\lambda_n \|\hat{\omega}^\top \beta\|_1$ to nullify "excessive" coefficients; λ_n and $\hat{\omega}$ controlling penalization; constraining $\beta \leq 0$ yields ALQR [▶ Details](#)

$$\hat{\beta}_{\tau, \lambda_n}^{\text{adapt}} = \arg \min_{\beta \in \mathbb{R}^p} \sum_{i=1}^n \rho_\tau(Y_i - X_i^\top \beta) + \lambda_n \|\hat{\omega}^\top \beta\|_1 \quad (1)$$



TEDAS Step 1

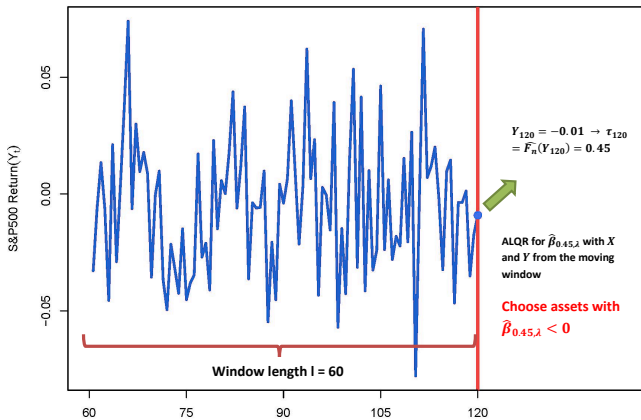
Initial wealth $W_0 = \$1$, $t = l, \dots, n$; $l = 60$ length of the moving window

□ Portfolio constituents' selection

1. determine core asset return Y_t , set $\tau_t = \widehat{F}_n(Y_t)$ ▶ Notation
2. ALQR for $\widehat{\beta}_{\tau_t, \lambda_n}$ using the observations $X \in \mathbb{R}^{t-l+1, \dots, t \times p}$, $Y \in \mathbb{R}^{t-l+1, \dots, t}$
3. if $Y_t < 0$, choose X_j , $j = 1, \dots, k < p$ with $\widehat{\beta}_{\tau_t, \lambda_n} < 0$; if $Y_t > 0$, choose X_j , $j = 1, \dots, k < p$ with $\widehat{\beta}_{\tau_t, \lambda_n} > 0$



TEDAS Step 1



TEDAS Step 2

□ Portfolio selection

1. apply TEDAS Gestalt to X_j , obtain $\hat{w}_t \in \mathbb{R}^k$
2. determine the realized portfolio wealth for $t + 1$,
$$\hat{X}_{t+1} \stackrel{\text{def}}{=} (X_{t+1,1}, \dots, X_{t+1,k})^\top: W_{t+1} = W_t(1 + \hat{w}_t^\top \hat{X}_{t+1})$$



TEDAS Example

1. Suppose $t = 86$ (Feb. 2007), $W_{86} = \$1.125$, accumulated wealth $W_{86} = \$1.429$, $Y_{86} = -1.85\% < 0$
2. $\hat{F}_n(Y_{86}) = 0.25$, so estimate $\hat{\beta}_{0.25} < 0$
3. ALQR on $X \in \mathbb{R}^{60 \times 163}$, $Y \in \mathbb{R}^{60}$ yields $\hat{\beta}_{0.25} = (-0.77, -1.12, -0.41)^\top$, *Latin American Arbitrage, North America Macro, Emerging Markets CTA/Managed Futures*
4. TEDAS CF-CVaR optimization [Details](#) yields $\hat{w}_{86} = (0.22, 0.16, 0.62)^\top$; $\hat{X}_{87} = (0.38\%, 0.45\%, 0.76\%)^\top$, $W_{87} = W_{86}(1 + \hat{w}_{86}^\top \hat{X}_{87}) = \1.438 (return of 0.62%) while $Y_{87} = -1.53\%$



TEDAS Gestalten

| TEDAS gestalt | Dynamics modelling | Weights optimization |
|----------------|---|---|
| TEDAS Naïve | NO | Equal weights |
| TEDAS Hybrid | NO | Mean-variance optimization of weights ▶ Details |
| TEDAS Basic | DCC volatility ▶ Details | CF-VaR optimization ▶ Details |
| TEDAS Advanced | Time-Varying ▶ Details | Cornish-Fisher-CVaR minimization ▶ Details |
| TEDAS Expert | Conditional Distributions | Expected utility optimization ▶ Details |



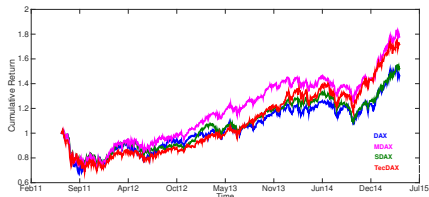
Hedge funds

- Monthly data
 - ▶ Core asset (Y): S&P 500, Nikkei225, DAX 30, FTSE 100
 - ▶ Satellite assets (X): 164 Eureka hedge hedge funds indices
- Span: 20000131-20131031 (166 months)
- Source: Bloomberg



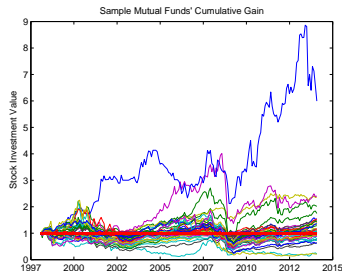
German equity

- Frankfurt Stock Exchange (Xetra), weekly data
 - ▶ Core asset (Y): DAX index
 - ▶ Satellites assets (X): 122 stocks - SDAX (48), MDAX (45) and TecDAX (49) as on 20140801
- Span: 20110707 - 20150424 (197 weeks)
- Source: Datastream



Mutual Funds

- Monthly data
 - ▶ Core asset (Y): S&P500
 - ▶ Satellite assets (X): 583 Mutual funds
- Span: 19980101 - 20131231
- Source: Datastream



Benchmark Strategies

1. **RR**: dynamic risk-return optimization [▶ Details](#)
2. **PES**: tail risk optimization [▶ Details](#)
3. **Risk-parity portfolio** (equal risk contribution) [▶ Details](#)
4. **60/40 portfolio** [▶ Details](#)



TEDAS with $Y = \text{S\&P 500}$

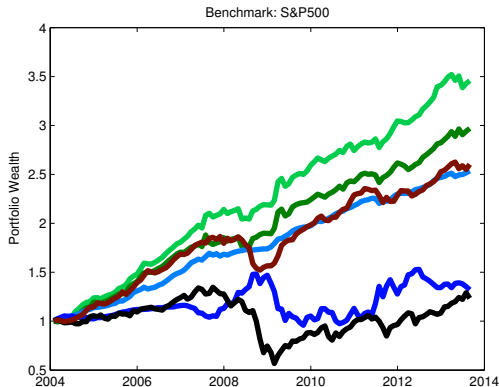


Figure 5: Cumulative portfolio wealth comparison: TEDAS Naïve, TEDAS Expert, TEDAS Advanced, RR, PESS, S&P 500 buy & hold; X = hedge funds' indices' returns matrix



TEDAS with $Y = \text{Nikkei 225}$

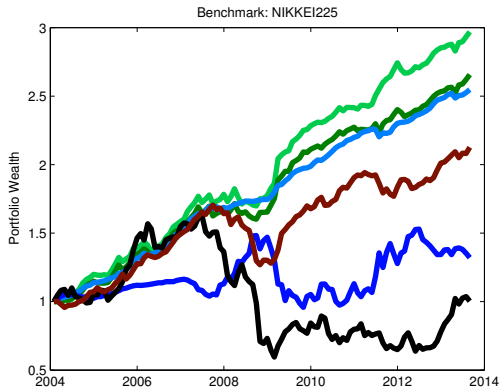


Figure 6: Cumulative portfolio wealth comparison: TEDAS Naïve, TEDAS Expert, TEDAS Advanced, RR, PESS, Nikkei 225 buy & hold; X = hedge funds' indices' returns matrix



TEDAS with $Y = \text{FTSE 100}$

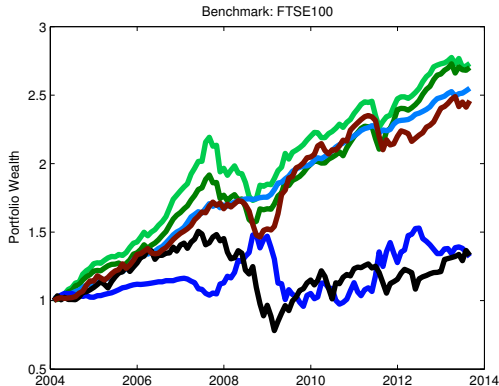


Figure 7: Cumulative portfolio wealth comparison: TEDAS Naïve, TEDAS Expert, TEDAS Advanced, RR, PESS, FTSE100 buy & hold; X = hedge funds' indices' returns matrix



TEDAS with $Y = \text{DAX30}$

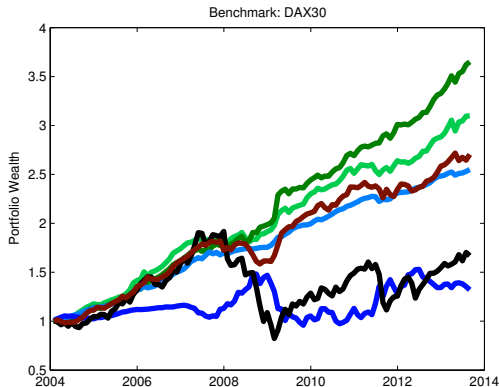


Figure 8: Cumulative portfolio wealth comparison: TEDAS Naïve, TEDAS Expert, TEDAS Advanced, RR, PESS, DAX30 buy & hold; X = hedge funds' indices' returns matrix



Histograms of \hat{q}

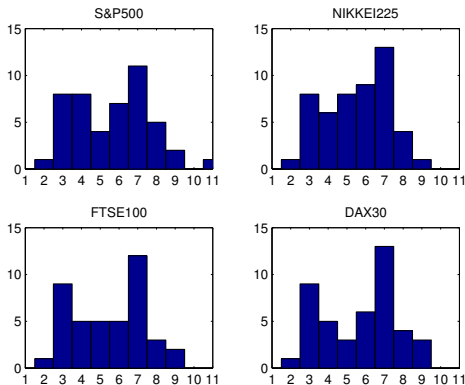


Figure 9: Frequency of the number of selected variables for 4 different Y



Selected Hedge Funds: S&P 500

Table 2: The selected hedge funds for S&P 500 benchmark

| Top 5 influential hedge funds | Frequency |
|---|-----------|
| Latin American Onshore Fixed Income Hedge Fund Index | 19 |
| Emerging Markets Dual Approach Absolute Return Fund Index | 18 |
| Large North American Hedge Fund Index | 15 |
| Europe Macro Hedge Fund Index | 15 |
| Asia Macro Hedge Fund Index | 14 |



Selected Hedge Funds: Nikkei 225

Table 3: The selected hedge funds for Nikkei 225 benchmark

| Top 5 influential hedge funds | Frequency |
|--|-----------|
| Small Latin American Hedge Fund Index | 22 |
| Taiwan Hedge Fund Index | 20 |
| North America Top-Down Absolute Return Fund Index | 16 |
| Large North American Hedge Fund Index | 15 |
| Latin American Onshore Fixed Income Hedge Fund Index | 15 |



Selected Hedge Funds: FTSE 100

Table 4: The selected hedge funds for FTSE 100 benchmark

| Top 5 influential hedge funds | Frequency |
|--|-----------|
| Small Latin American Hedge Fund Index | 15 |
| Europe Macro Hedge Fund Index | 15 |
| Latin American Onshore Fixed Income Hedge Fund Index | 14 |
| Asia Pacific Top-Down Absolute Return Fund Index | 14 |
| Latin American Fixed Income Hedge Fund Index | 13 |



Selected Hedge Funds: DAX 30

Table 5: The selected hedge funds for DAX 30 benchmark

| Top 5 influential hedge funds | Frequency |
|---|-----------|
| Emerging Markets Dual Approach Absolute Return Fund Index | 21 |
| North America Macro Hedge Fund Index | 19 |
| Taiwan Hedge Fund Index | 16 |
| Asia CTA Hedge Fund Index | 14 |
| Europe Macro Hedge Fund Index | 14 |



Dynamic Moment Parameters

Table 6: Parameter estimates in (14) and (15) at $t = 160$

| | s_1 | s_2 | s_3 | s_4 | s_5 | s_6 |
|---------------|----------|----------|---------|----------|---------|----------|
| ω | 0.12*** | 0.33*** | 0.40*** | 0.61*** | 0.60*** | 0.47*** |
| α_1 | 0.72*** | 0.29*** | 0.47*** | 0.30*** | 0.33*** | 0.45*** |
| β_1 | 0.13*** | 0.35*** | 0.12*** | 0.05** | 0.07*** | 0.06 |
| a_0 | 0.05 | -0.02 | -0.27 | 0.08*** | 0.11 | -0.08 |
| a_1^- | -0.33* | 0.08 | -0.55 | -0.13*** | -0.13 | -0.69** |
| a_1^+ | -0.43 | 0.22*** | 1.17 | -0.46*** | -0.33 | -0.47*** |
| a_2 | 1.45*** | 2.16*** | 0.96 | 2.14*** | 1.77*** | 1.10** |
| b_0 | -5.03*** | -1.96* | -5.53 | -3.65*** | -4.16** | -1.89*** |
| b_1^- | 0.70** | 0.21 | 0.88 | 1.56 | -1.53 | 0.51*** |
| b_1^+ | 0.79*** | -1.40*** | 0.87 | 0.89 | -1.87 | -0.79*** |
| b_2 | -11.67 | -0.69 | 0.12 | -0.64*** | 0.38 | 1.11*** |
| $Factor_{LL}$ | -193.87 | -201.40 | -215.14 | -216.53 | -218.57 | -225.10 |
| $Model_{LL}$ | 2394 | | | | | |

parameter estimates under the *NIG* distribution, for the log-returns of Eurekahedge 6 hedge funds (160 data monthly returns, 01.01.2000 - 30.04.2013)

the conditional variance of factors follows a GARCH(1,1) model

the conditional dynamics of skew and shape parameters is bounded by a logistic transformation

the *, ** and *** denote significance at the 10%, 5% and 1% levels, respectively



Evolution of ξ_t and ν_t

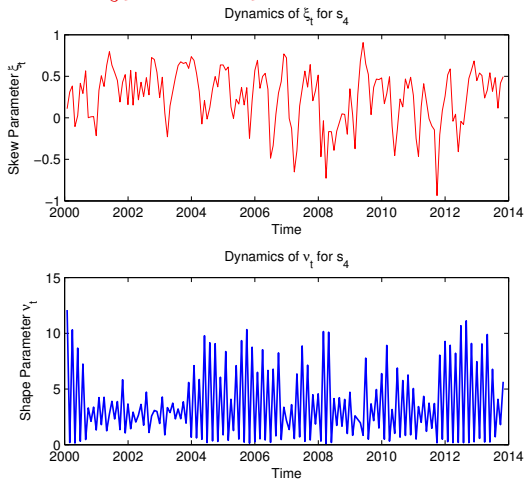


Figure 10: Evolution of the skew and shape parameters for s_4 in Table 6



Conditional Skewness and Kurtosis

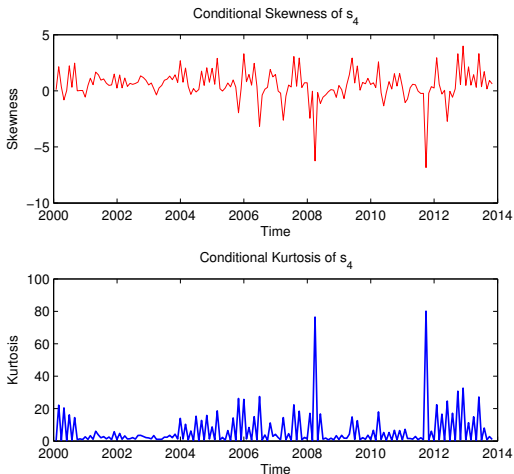


Figure 11: Evolution of the conditional skewness and kurtosis for s_4



TEDAS: DAX results

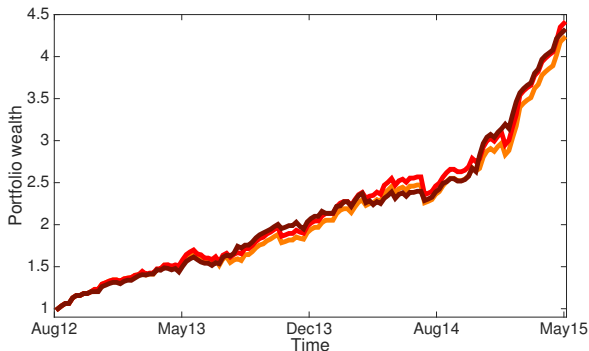



Figure 12: Strategies' cumulative returns' comparison (with transaction costs 1% of portfolio value): TEDAS Naïve, TEDAS Hybrid, TEDAS basic
 TEDAS_strategies



TEDAS: DAX results

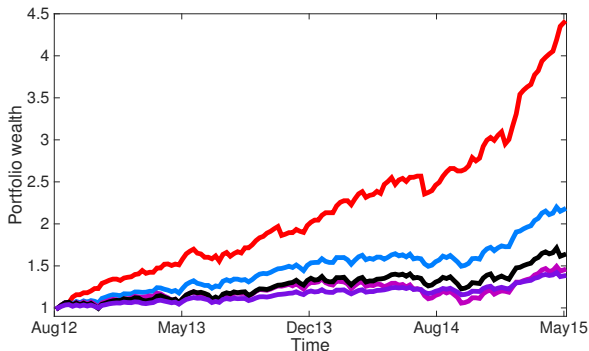



Figure 13: Strategies' cumulative returns' comparison (with transaction costs 1% of portfolio value): TEDAS Hybrid, DAX Buy-and-hold, 60/40, Risk-parity, RR

 TEDAS_strategies



TEDAS: Mutual Funds' results

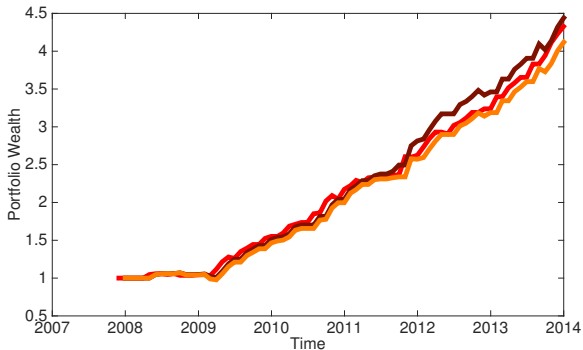



Figure 14: Strategies' cumulative returns' comparison (with transaction costs 1% of portfolio value): TEDAS Naïve, TEDAS Hybrid, TEDAS basic
 TEDAS_strategies



TEDAS: Mutual Funds' results

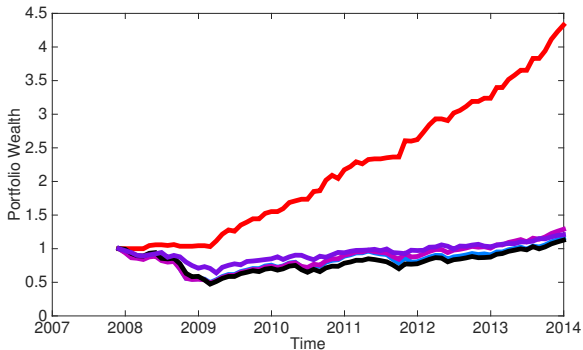


Figure 15: Strategies' cumulative returns' comparison (with transaction costs 1% of portfolio value): TEDAS Hybrid, DAX Buy-and-hold, 60/40, Risk-parity, RR

 TEDAS_strategies



Conclusions: TLND challenge

- Lasso quantile regression captures **T**ail events
- Lasso asset selection resolves **L**arge dimensionality problem
- Higher-moment optimization for **N**on-normality
- **D**ynamic portfolio optimization through conditional distribution modelling
- TEDAS: out-of-sample performance superior to benchmark strategies



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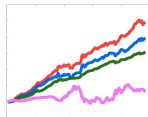
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Notation

□ $\hat{q}_\tau \stackrel{\text{def}}{=} \hat{F}_n^{-1}(\tau)$, with

$$\hat{F}_n(Y_t) \stackrel{\text{def}}{=} \int_{-\infty}^{Y_t} \hat{f}_n(u) du = \frac{1}{n} \sum_{i=1}^n H\left(\frac{Y_t - Y_i}{h}\right), \quad (2)$$

where $\hat{f}_n(Y_t) \stackrel{\text{def}}{=} (1/nh) \sum_{i=1}^n K\{(Y_t - Y_i)/h\}$,
 $H(x) = \int_{-\infty}^x K(u) du$, $K(\cdot) = \varphi(\cdot)$;

Silverman (1986) rule-of-thumb:

$$h = 1.06sn^{-1/5}, \quad s \text{ sample standard deviation of } Y$$

□ $\hat{\beta}_{\tau, \lambda_n}$ are the estimated non-zero ALQR coefficients

▶ Back to "TEDAS Step 1"



Lasso Shrinkage

Linear model: $Y = X\beta + \varepsilon$; $Y \in \mathbb{R}^n$, $X \in \mathbb{R}^{n \times p}$, $\beta \in \mathbb{R}^p$, $\{\varepsilon_i\}_{i=1}^n$ i.i.d., independent of $\{X_i; i = 1, \dots, n\}$

The optimization problem for the lasso estimator:

$$\begin{aligned} \hat{\beta}^{\text{lasso}} &= \arg \min_{\beta \in \mathbb{R}^p} f(\beta) \\ &\text{subject to } g(\beta) \geq 0 \end{aligned} \tag{3}$$

where

$$\begin{aligned} f(\beta) &= \frac{1}{2} (y - X\beta)^\top (y - X\beta) \\ g(\beta) &= t - \|\beta\|_1 \end{aligned}$$

where t is the size constraint on $\|\beta\|_1$ [▶ Back to "Tail Events"](#)



Lasso Duality

If (1) is convex programming problem, then the Lagrangian is

$$L(\beta, \lambda) = f(\beta) - \lambda g(\beta).$$

and the primal-dual relationship is

$$\underbrace{\text{minimize}_{\beta} \sup_{\lambda \geq 0} L(\beta, \lambda)}_{\text{primal}} \geq \underbrace{\text{maximize}_{\lambda \geq 0} \inf_{\beta} L(\beta, \lambda)}_{\text{dual}}$$

Then the dual function $L^*(\lambda) = \inf_{\beta} L(\beta, \lambda)$ is

$$L^*(\lambda) = \frac{1}{2} y^T y - \frac{1}{2} \hat{\beta}^T X^T X \hat{\beta} - t \frac{(y - X \hat{\beta})^T X \hat{\beta}}{\|\hat{\beta}\|_1}$$

with $(y - X \hat{\beta})^T X \hat{\beta} / \|\hat{\beta}\|_1 = \lambda$ [▶ Back to "Tail Events"](#)



Paths of Lasso Coefficients

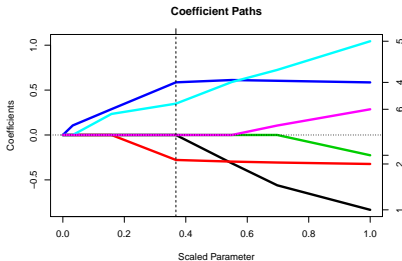


Figure 16: Lasso shrinkage of coefficients in the hedge funds dataset example (6 covariates were chosen for illustration); each curve represents a coefficient as a function of the scaled parameter $\hat{s} = t/\|\beta\|_1$; the dashed line represents the model selected by the BIC information criterion ($\hat{s} = 3.7$)

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Example of Lasso Geometry

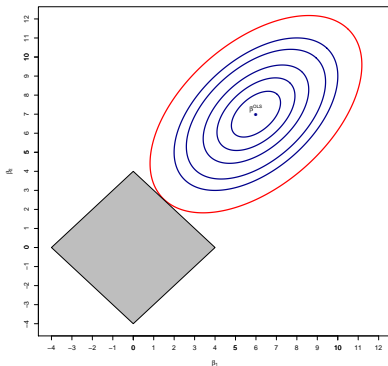



Figure 17: Contour plot of the residual sum of squares objective function centered at the OLS estimate $\hat{\beta}^{ols} = (6, 7)$ and the constraint region $\sum |\beta_j| \leq t$  MVAlassocontour



Quantile Regression

The loss $\rho_\tau(u) = u\{\tau - \mathbf{I}(u < 0)\}$ gives the (conditional) quantiles $F_{y|x}^{-1}(\tau) \stackrel{\text{def}}{=} q_\tau(x)$.

Minimize

$$\hat{\beta}_\tau = \arg \min_{\beta \in \mathbb{R}^p} \sum_{i=1}^n \rho_\tau(Y_i - X_i^\top \beta).$$

Re-write:

$$\underset{(\xi, \zeta) \in \mathbb{R}_+^{2n}}{\text{minimize}} \quad \left\{ \tau \mathbf{1}_n^\top \xi + (1 - \tau) \mathbf{1}_n^\top \zeta \mid X\beta + \xi - \zeta = Y \right\}$$

with ξ, ζ are vectors of "slack" variables [▶ Back to "Tail Events"](#)



Non-Positive (NP) Lasso-Penalized QR

The **lasso-penalized** QR problem with an additional non-positivity constraint takes the following form:

$$\begin{aligned} & \underset{(\xi, \zeta, \eta, \tilde{\beta}) \in \mathbb{R}_+^{2n+p} \times \mathbb{R}^p}{\text{minimize}} && \tau \mathbf{1}_n^\top \xi + (1 - \tau) \mathbf{1}_n^\top \zeta + \lambda \mathbf{1}_n^\top \eta \\ & \text{subject to} && \xi - \zeta = Y + X \tilde{\beta}, \\ & && \xi \geq \mathbf{0}, \\ & && \zeta \geq \mathbf{0}, \\ & && \eta \geq \tilde{\beta}, \\ & && \eta \geq -\tilde{\beta}, \\ & && \tilde{\beta} \geq \mathbf{0}, \quad \tilde{\beta} \stackrel{\text{def}}{=} -\beta \end{aligned} \tag{4}$$

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Solution

Transform into matrix (I_p is $p \times p$ identity matrix; $E_{p \times n} = \begin{pmatrix} I_n \\ 0 \end{pmatrix}$):

$$\begin{aligned} & \text{minimize} && c^\top x \\ & \text{subject to} && Ax = b, \quad Bx \leq 0 \end{aligned}$$

where $A = \begin{pmatrix} I_n & -I_n & 0 & X \end{pmatrix}$, $b = Y$, $x = \begin{pmatrix} \xi & \zeta & \eta & \beta \end{pmatrix}^\top$,

$$c = \begin{pmatrix} \tau 1_n \\ (1 - \tau) 1_n \\ \lambda 1_p \\ 0 1_p \end{pmatrix}, \quad B = \begin{pmatrix} -E_{p \times n} & 0 & 0 & 0 \\ 0 & -E_{p \times n} & 0 & 0 \\ 0 & 0 & -I_p & I_p \\ 0 & 0 & -I_p & -I_p \\ 0 & 0 & 0 & I_p \end{pmatrix}$$

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Solution - Continued

The previous problem may be reformulated into *standard form*

$$\begin{aligned} & \text{minimize} && c^\top x \\ & \text{subject to} && Cx = d, \\ & && x + s = u, \quad x \geq 0, s \geq 0 \end{aligned}$$

and the dual problem is:

$$\begin{aligned} & \text{maximize} && d^\top y - u^\top w \\ & \text{subject to} && C^\top y - w + z = c, \quad z \geq 0, w \geq 0 \end{aligned}$$

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Solution - Continued

The KKT conditions for this linear program are

$$F(x, y, z, s, w) = \begin{Bmatrix} Cx - d \\ x + s - u \\ C^T y - w + z - c \\ x \circ z \\ s \circ w \end{Bmatrix} = 0,$$

with $y \geq 0$, $z \geq 0$ dual slacks, $s \geq 0$ primal slacks, $w \geq 0$ dual variables.

This can be solved by a primal-dual path following algorithm based on the *Newton method*

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Adaptive Lasso Procedure

Lasso estimates $\hat{\beta}$ can be inconsistent (Zou, 2006) in some scenarios.

Lasso soft-threshold function gives biased results

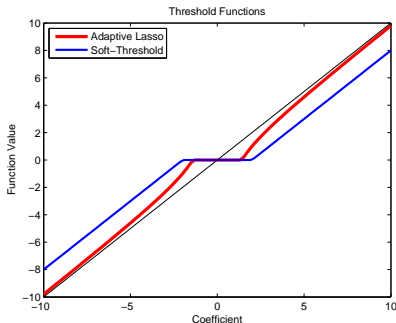


Figure 18: Threshold functions for simple and adaptive Lasso



Adaptive Lasso Procedure

The adaptive Lasso (Zou, 2006) yields a sparser solution and is less biased.

L_1 - penalty replaced by a re-weighted version; $\hat{\omega}_j = 1/|\hat{\beta}_j^{\text{init}}|^\gamma$,
 $\gamma = 1$, $\hat{\beta}^{\text{init}}$ is from (3)

The adaptive lasso estimates are given by:

$$\hat{\beta}_\lambda^{\text{adapt}} = \arg \min_{\beta \in \mathbb{R}^p} \sum_{i=1}^n (Y_i - X_i^\top \beta)^2 + \lambda \|\hat{\omega}^\top \beta\|_1$$

(Bühlmann, van de Geer, 2011): $\hat{\beta}_j^{\text{init}} = 0$, then $\hat{\beta}_j^{\text{adapt}} = 0$

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Simple and Adaptive Lasso Penalized QR

Simple lasso-penalized QR optimization problem is:

$$\hat{\beta}_{\tau,\lambda} = \arg \min_{\beta \in \mathbb{R}^p} \sum_{i=1}^n \rho_{\tau}(Y_i - X_i^{\top} \beta) + \lambda \|\beta\|_1 \quad (5)$$

Adaptive lasso-penalized QR model uses the re-weighted penalty:

$$\hat{\beta}_{\tau,\lambda}^{\text{adapt}} = \arg \min_{\beta \in \mathbb{R}^p} \sum_{i=1}^n \rho_{\tau}(Y_i - X_i^{\top} \beta) + \lambda \|\hat{\omega}^{\top} \beta\|_1 \quad (6)$$

Adaptive lasso-penalized QR procedure can ensure oracle properties for the estimator [▶ Details](#)

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Algorithm for Adaptive Lasso Penalized QR

The optimization for the adaptive lasso quantile regression can be re-formulated as a lasso problem:

- the covariates are rescaled: $\tilde{X} = (X_1 \circ \hat{\beta}_1^{\text{init}}, \dots, X_p \circ \hat{\beta}_p^{\text{init}})$;
- the lasso problem (5) is solved:

$$\hat{\beta}_{\tau, \lambda} = \arg \min_{\beta \in \mathbb{R}^p} \sum_{i=1}^n \rho_{\tau}(Y_i - \tilde{X}_i^{\top} \beta) + \lambda \|\beta\|_1$$

- the coefficients are re-weighted as $\hat{\beta}_{\tau, \lambda}^{\text{adapt}} = \hat{\beta}_{\tau, \lambda} \circ \hat{\beta}^{\text{init}}$

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Oracle Properties of an Estimator

An estimator has oracle properties if (Zheng et al., 2013):

- it selects the correct model with probability converging to 1;
- the model estimates are consistent with an appropriate convergence rate;
- estimates are asymptotically normal with the same asymptotic variance as that knowing the true model

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Oracle Properties for Adaptive Lasso QR

In the linear model, let $Y = X\beta + \varepsilon = X^1\beta^1 + X^2\beta^2 + \varepsilon$, where $X = (X^1, X^2)$, $X^1 \in \mathbb{R}^{n \times q}$, $X^2 \in \mathbb{R}^{n \times (p-q)}$; β_q^1 are true nonzero coefficients, $\beta_{p-q}^2 = 0$ are noise coefficients; $q = \|\beta\|_0$.

Also assume that $\lambda q / \sqrt{n} \rightarrow 0$ and $\lambda / \{\sqrt{q} \log(n \vee p)\} \rightarrow \infty$ and certain regularity conditions are satisfied [▶ Details](#)

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Oracle Properties for Adaptive Lasso QR

Then the adaptive L_1 QR estimator has the oracle properties (Zheng et al., 2013):

1. Variable selection consistency:

$$P(\beta^2 = 0) \geq 1 - 6 \exp \left\{ -\frac{\log(n \vee p)}{4} \right\}.$$

2. Estimation consistency: $\|\beta - \hat{\beta}\| = \mathcal{O}_p(\sqrt{q/n})$
3. Asymptotic normality: $u_q^2 \stackrel{\text{def}}{=} \alpha^T \Sigma_{11} \alpha, \forall \alpha \in \mathbb{R}^q, \|\alpha\| < \infty,$

$$n^{1/2} u_q^{-1} \alpha^T (\beta^1 - \hat{\beta}^1) \xrightarrow{\mathcal{L}} N \left\{ 0, \frac{(1-\tau)\tau}{f^2(\gamma^*)} \right\}$$

where γ^* is the τ th quantile and f is the pdf of ε

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Selected Hedge Funds' Strategies

1. *Convertible arbitrage* hedge funds focus on the mispricing of convertible bonds. A typical position involves a long position in the convertible bond and a short position in the underlying asset. N
2. *Fixed income arbitrage* hedge funds tend to profit from price anomalies between related securities and/or bet on the evolution of interest rates spreads. Typical trading strategies are butterfly-like structures, cash/futures basis trading strategies or relative swap spread trades. D
3. *Event-driven* hedge funds focus on price movements generated by an anticipated corporate event, such as a merger, an acquisition, a bankruptcy, etc. T

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Selected Hedge Funds' Strategies

4. *Long/short equity* hedge funds represent the original hedge fund model. They invest in equities both on the long and the short sides, and generally have a small net long exposure. They are genuinely opportunistic strategies and could be classified as "double alpha, low beta" funds. 
5. *Market neutral* hedge funds seek to neutralize certain market risks by taking offsetting long and short positions in instruments with actual or theoretical relationships. Most of them are in fact long/short equity hedge funds.
6. *Dedicated short bias* hedge funds are essentially long/short equity hedge funds, that maintain a consistent net short exposure, therefore attempting to capture profits when the market declines. [▶ Back](#)



Selected Hedge Funds' Strategies

7. *Emerging market* hedge funds invest in equities and fixed-income securities of emerging markets around the world.
8. *Global macro* hedge funds take very large directional bets on overall market directions that reflect their forecasts of major economic trends and/or events.
9. *Managed futures* hedge funds implement discretionary or systematic trading in listed financial, commodity and currency futures around the world. The managers of these funds are known as commodity trading advisors (CTAs).
10. *Multi-strategy* hedge funds regroup managers acting in several of the above-mentioned strategies. [▶ Return](#)



Traditional Assets/Hedge Fund Indices

Table 7: Correlation statistics for MSCI and hedge funds' indices returns

| Hedge Fund Indices | MSCI Indices | | | | | | | | |
|----------------------------|--------------|-------|-------|-------|-------|-------|-------|-------|------|
| | WRD | EUR | US | UK | FR | SW | GER | JAP | PAC |
| Asia CTA | -0.01 | 0.02 | -0.02 | -0.06 | 0.01 | -0.09 | 0.04 | -0.03 | 0.02 |
| Asia Distressed Debt | 0.30 | 0.30 | 0.24 | 0.31 | 0.31 | 0.26 | 0.27 | 0.26 | 0.34 |
| Asia Macro | -0.01 | -0.01 | -0.04 | 0.01 | -0.02 | 0.07 | -0.03 | 0.06 | 0.06 |
| Global CTA FoF | 0.02 | 0.08 | -0.08 | 0.09 | 0.10 | 0.09 | 0.07 | 0.06 | 0.10 |
| Global Event Driven FoF | 0.65 | 0.59 | 0.58 | 0.66 | 0.59 | 0.50 | 0.57 | 0.47 | 0.67 |
| Global Macro FoF | 0.19 | 0.22 | 0.07 | 0.24 | 0.22 | 0.18 | 0.20 | 0.23 | 0.31 |
| CTA/Managed Futures | -0.04 | 0.02 | -0.13 | 0.03 | 0.03 | 0.07 | -0.01 | 0.04 | 0.05 |
| Event Driven | 0.82 | 0.75 | 0.75 | 0.78 | 0.75 | 0.64 | 0.75 | 0.62 | 0.83 |
| Fixed Income | 0.70 | 0.65 | 0.63 | 0.70 | 0.65 | 0.56 | 0.62 | 0.51 | 0.78 |
| Long Short Equities | 0.82 | 0.78 | 0.74 | 0.76 | 0.77 | 0.64 | 0.77 | 0.64 | 0.82 |
| Asia inc Japan Distr. Debt | 0.30 | 0.30 | 0.24 | 0.31 | 0.31 | 0.26 | 0.27 | 0.26 | 0.34 |
| Asia inc Japan Macro | 0.34 | 0.33 | 0.31 | 0.27 | 0.33 | 0.24 | 0.35 | 0.31 | 0.40 |

Calculations based on monthly data Jan. 2000 - Jul. 2012

WRD - World, EUR - Eurozone, FR - France, SW - Switzerland, PAC - Pacific ex. Japan

FoF means "fund of funds"

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Risk-Return Asset Allocation

Log returns $X_t \in \mathbb{R}^p$:

$$\begin{aligned} \min_{w_t \in \mathbb{R}^p} \quad & \sigma_{P,t}^2(w_t) \stackrel{\text{def}}{=} w_t^\top \Sigma_t w_t \\ \text{s.t.} \quad & \mu_{P,t}(w_t) = r_T, \\ & w_t^\top \mathbf{1}_p = 1, \\ & w_{i,t} \geq 0 \end{aligned} \tag{7}$$

where r_T "target" return, $\Sigma_t \stackrel{\text{def}}{=} E_{t-1}\{(X_t - \mu)(X_t - \mu)^\top\}$, Σ_t is modeled with a GARCH model

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Tail Risk Asset Allocation

Portfolio returns $X \in \mathbb{R}^{n \times p}$, Bassett et al. (2004)

$$\begin{aligned} \min_{(\beta, \alpha)^T \in \mathbb{R}^p} \quad & \sum_{t=1}^n \rho_{\tau} \left\{ X_{t1} - \sum_{j=2}^p (X_{t1} - X_{tj}) \beta_j - \alpha \right\} \\ \text{s.t.} \quad & w^{\top} \hat{\mu} = r_{\mathcal{T}}, \\ & w^{\top} \mathbf{1}_p = 1, \end{aligned} \quad (8)$$

where $r_{\mathcal{T}}$ is the "target" return for the portfolio and $w = w(\beta) = (1 - \sum_{j=2}^p \beta_j, \beta^{\top})^{\top}$, $\tau \in (0, 1)$, $\hat{\mu} \stackrel{\text{def}}{=} \bar{X}$ sample returns' mean

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Multi-Moment Utility Optimization

The (dynamic) investment decision: $U(\cdot)$ utility function; $X_t \in \mathbb{R}^p$ log-returns, w_t weights, $\mu_{P,t}(w_t) \stackrel{\text{def}}{=} w_t^\top \mu$, $\mu \stackrel{\text{def}}{=} E_{t-1}(X_t)$, r_T "target" return:

$$\max_{w_t \in \mathbb{R}^p} E_{t-1} \{U(W_t)\}, \quad \text{s.t. } \mu_{P,t}(w_t) = r_T, \quad w^\top \mathbf{1}_p = 1, \quad w_{i,t} \geq 0, \quad (9)$$

$$\begin{aligned} E_{t-1} \{U(W_t)\} &\approx U\{\bar{W}_t\} + \frac{1}{2} U^{(2)}\{\bar{W}_t\} \sigma_{W_t}^2 + \\ &+ \frac{1}{3!} U^{(3)}\{\bar{W}_t\} S_{W_t} + \frac{1}{4!} U^{(4)}\{\bar{W}_t\} K_{W_t}, \end{aligned}$$

where $W_t \stackrel{\text{def}}{=} 1 + w_t^\top X_t$ is the end-of-period t wealth, $\bar{W}_t \stackrel{\text{def}}{=} E_{t-1}(W_t)$, $\sigma_{W_t}^2 \stackrel{\text{def}}{=} E_{t-1} \{(W_t - \bar{W}_t)^2\}$, $S_{W_t} \stackrel{\text{def}}{=} E_{t-1} \{(W_t - \bar{W}_t)^3\}$, $K_{W_t} \stackrel{\text{def}}{=} E_{t-1} \{(W_t - \bar{W}_t)^4\}$; $U^{(n)}(\cdot)$ is the n th derivative of $U(\cdot)$

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Utility Function Example

- CARA utility:

$$U(W) = -\exp(-\eta W),$$

where η coefficient of risk aversion

- then:

$$\begin{aligned} E_{t-1} \{U(W_t)\} &= E_{t-1} \{-\exp(-\eta W_t)\} \\ &\approx -\exp(-\eta \bar{W}_t) \left(1 + \frac{\eta^2}{2} \sigma_{W_t}^2 - \frac{\eta^3}{3!} S_{W_t} + \frac{\eta^4}{4!} K_{W_t} \right) \end{aligned}$$

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Portfolio Moments

The portfolio moments:

$$\sigma_{W_t}^2 = w_t^\top M_t^2 w_t$$

$$S_{W_t} = w_t^\top M_t^3 (w_t \otimes w_t)$$

$$K_{W_t} = w_t^\top M_t^4 (w_t \otimes w_t \otimes w_t),$$

where \otimes Kronecker product,

$$M_t^2 \stackrel{\text{def}}{=} E_{t-1}(r_t - \mu)^2 \quad (10)$$

$$M_t^3 \stackrel{\text{def}}{=} E_{t-1}\{(r_t - \mu)(r_t - \mu)^\top \otimes (r_t - \mu)^\top\} \quad (11)$$

$$M_t^4 \stackrel{\text{def}}{=} E_{t-1}\{(r_t - \mu)(r_t - \mu)^\top \otimes (r_t - \mu)^\top \otimes (r_t - \mu)^\top\}, \quad (12)$$

A dynamic distribution model is used to obtain M_t^2 , M_t^3 , M_t^4 in (10), (11), (12) [▶ Return to "TEDAS Gestalten"](#)



Dynamic Distribution Model

- joint normality is questionable
- possible persistence in the dynamics of moments
- reaction of distribution parameters to past shocks
- computational feasibility

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Descriptive Statistics

Table 8: Monthly returns of 3 Eureka hedge funds' indices

| | Japan Multi-Strategy | | North America Fixed Income | | Europe Arbitrage | |
|------------------------------------|----------------------|---------|----------------------------|---------|------------------|---------|
| Univariate statistics | | | | | | |
| Normality tests | | | | | | |
| JB | 533.775 | (0.000) | 294.089 | (0.000) | 610.407 | (0.000) |
| KS | 0.503 | (0.000) | 0.473 | (0.000) | 0.485 | (0.000) |
| Omnibus | 82.773 | (0.000) | 43.761 | (0.000) | 171.079 | (0.000) |
| Dynamic conditional moments' tests | | | | | | |
| ARCH | 11.227 | (0.000) | 34.966 | (0.000) | 26.592 | (0.000) |
| Bera-Lee | 48.469 | (0.000) | 36.475 | (0.000) | 40.783 | (0.000) |
| Bera-Zuo | 203.723 | (0.000) | 20.149 | (0.166) | 421.847 | (0.000) |
| Multivariate statistics | | | | | | |
| Test | | | | | | |
| Omnibus | 326.226 | (0.000) | | | | |
| Mardia | 301.199 | (0.000) | | | | |
| Henze-Zirkler | 9.862 | (0.000) | | | | |

Standard errors and p -values are given in parentheses.

ARCH, Bera-Lee and Bera-Zuo stand for the test statistics of the ARCH test by Engle (1982) and information matrix tests for testing variation in second, third and fourth conditional moments

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Generalized Hyperbolic (GH) Distribution

A vector X has a multivariate GH distribution if

$$X = \mu + W\delta + \sqrt{W}AZ, \quad (13)$$

where

- (i) $Z \sim N(0, I_k)$
- (ii) $A \in \mathbb{R}^{d \times k}$
- (iii) $\mu, \delta \in \mathbb{R}^d$
- (iv) $W \geq 0$, scalar-valued random variable, independent of Z ,
 $W \sim GIG(\lambda, \alpha, \beta)$; GIG is the generalized inverse Gaussian distribution

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Multivariate Affine GH Distribution

- margins of the (MGH) distribution not mutually independent for some choice of $\Sigma = AA^T$
- MAGH distribution, Schmidt et al. (2006), models margins and dependency independently

$Y \sim \text{MAGH}(\lambda, \alpha, \beta, \mu, \Sigma)$ if

- $X = (X_1, \dots, X_d)^T$, $X_i \sim \text{GH}(0, 1, \alpha_i, \beta_i)$, $i = 1, \dots, d$
- $Y = AX + \mu$, $AA^T = \Sigma$ positive definite

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Normal Inverse Gaussian (NIG) Distribution

- obtained from the GH distribution with $\lambda = -0.5$
- "semi-heavy tails" property: fits financial data well

The density is written as:

$$f_{NIG}(x) = \frac{\alpha\delta}{\pi} \exp\left\{\delta\sqrt{\alpha^2 - \beta^2} + \beta(x - \mu)\right\} \frac{K_1\left\{\alpha\sqrt{\delta^2 + (x - \mu)^2}\right\}}{\sqrt{\delta^2 + (x - \mu)^2}},$$

where $0 \leq |\beta| \leq \alpha$, $\delta > 0$, K_1 is the modified Bessel function of the third kind and order 1

Location-Scale Property: let $\bar{\alpha} \stackrel{\text{def}}{=} \delta\alpha$ and $\bar{\beta} \stackrel{\text{def}}{=} \delta\beta$, then $X \sim NIG(\bar{\alpha}, \bar{\beta}, \mu, \delta) \Leftrightarrow (X - \mu)/\delta \sim NIG(\bar{\alpha}, \bar{\beta}, 0, 1)$

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Choice of the Matrix A

- assume $X = As$, X random signal generated by another random vector $s = (s_1, \dots, s_d)$, s_i statistically independent, $i = 1, \dots, d$ and a mixing matrix A , both unknown
- the *independent component analysis* (ICA) technique separates source signals s from a set of mixed signals X without or with very little aid of information about f or the mixing process A
- ICA estimates A and s by maximizing the nongaussianity of linear combinations of X

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The Model for Portfolio Returns

- assume $\varepsilon_t = A s_t$, $E \varepsilon_t = 0$, $E \varepsilon_t \varepsilon_t^\top = I_d$, $E s_t = 0$, $E s_t s_t^\top = I_d$
- define $E(s_t | \mathcal{F}_t) = 0$, $D_t \stackrel{\text{def}}{=} E(s_t s_t^\top | \mathcal{F}_t) \stackrel{\text{def}}{=} \text{diag}(d_{1t}, \dots, d_{dt})$
- let $z_{it} \sim NIG(\bar{\alpha}_{it}, \bar{\beta}_{it}, 0, 1)$, then $s_{it} \sim NIG(\bar{\alpha}_{it}/\sqrt{d_{it}}, \bar{\beta}_{it}/\sqrt{d_{it}}, 0, \sqrt{d_{it}})$
- *MANIG*: multivariate affine normal inverse Gaussian distribution
- model for portfolio returns $r_t = m_t + \varepsilon_t$,
 $r_t | \mathcal{F}_t \sim MANIG(m_t, \Sigma_t, \omega_t)$, where $\omega_t = (\omega_{1t}, \dots, \omega_{dt})^\top$ and $\omega_{it} = (\alpha_{it}, \beta_{it})^\top$, $i = 1, \dots, d$, $\Sigma_t = M_t^2 = A D_t A^\top$, d_{it} can be modeled as GARCH-type processes

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Moment Dynamics

- reparametrize the model to have asymmetry and shape parameters $\xi_{it} = \beta_{it}/\alpha_{it}$, $\nu_{it} = \sqrt{\alpha_{it}^2 - \beta_{it}^2}$
- introduce asymmetric GARCH-like dynamics:

$$\nu_{i,t} = a_{i,0} + a_{i,1}^- |s_{i,t-1}| N_{i,t-1} + a_{i,1}^+ |s_{i,t-1}| P_{i,t-1} + a_{i,2} \nu_{i,t-1} \quad (14)$$

$$\xi_{i,t} = b_{i,0} + b_{i,1}^- s_{i,t-1} N_{i,t-1} + b_{i,1}^+ s_{i,t-1} P_{i,t-1} + b_{i,2} \xi_{i,t-1}, \quad (15)$$

where $N_{i,t} = \mathbf{I}(z_{i,t} \leq 0)$, $P_{i,t} = 1 - N_{i,t}$

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Portfolio Moments

$$M_t^3 = AM_{S_t}^3(A \otimes A)^\top, \quad M_t^4 = AM_{S_t}^4(A \otimes A \otimes A)^\top,$$

where

$$M_{S_t}^3 = E_{t-1}(s_{i,t}s_{j,t}s_{k,t}) = \sum_{r=1}^p d_{ir,t}d_{jr,t}d_{kr,t}sk_{rt}^s$$

$$\begin{aligned} M_{S_t}^4 &= E_{t-1}(s_{i,t}s_{j,t}s_{k,t}s_{l,t}) \\ &= \sum_{r=1}^p d_{ir,t}d_{jr,t}d_{kr,t}d_{lr,t}kurt_{rt}^s + \sum_{r=1}^p \sum_{s \neq r} \psi_{rs,t}, \end{aligned}$$

$$\psi_{rs,t} = d_{ir,t}d_{jr,t}d_{ks,t}d_{ls,t} + d_{ir,t}d_{js,t}d_{kr,t}d_{ls,t} + d_{is,t}d_{jr,t}d_{kr,t}d_{ls,t},$$

$$D_t^{1/2} = (d_{ij,t})_{i,j=1,\dots,p}, \quad sk_{it}^s, \quad kurt_{it}^s \text{ are obtained with } \alpha_{it}, \beta_{it}$$

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Conditional VaR (CVaR) Optimization

Given $\alpha > 0.5$ confidence level,

$$\min_{w_t \in \mathbb{R}^p} \text{CVaR}_\alpha(w_t), \quad \text{s.t. } \mu_{P,t}(w_t) = r_T, w_t^\top \mathbf{1}_p = 1, w_{i,t} \geq 0, \quad (16)$$

$$\text{CVaR}_\alpha(w_t) = -\frac{1}{1-\alpha} q_{\alpha^*}^*(w_t) \sigma_{P,t}(w_t), \quad \text{▶ Proof} \quad (17)$$

where (via Cornish-Fisher (CF) expansion):

$$q_{\alpha^*}^*(w_t) = \left\{ 1 + \frac{S_{P,t}(w_t)}{6} z_{\alpha^*} + \frac{K_{P,t}(w_t)}{24} (z_{\alpha^*}^2 - 1) - \frac{S_{P,t}^2(w_t)}{36} (2z_{\alpha^*}^2 - 1) \right\} \varphi(z_{\alpha^*}), \quad (18)$$

where $\alpha^* \stackrel{\text{def}}{=} 1 - \alpha$ [▶ Return to "TEDAS Gestalten"](#) [▶ Back to "TEDAS Example"](#)



The Orthogonal GARCH Model

- $X_t \in \mathbb{R}^{n \times p}$, $\Gamma_t = B_t \in \mathbb{R}^{p \times p}$ matrix of standardized eigenvectors of $n^{-1}X_t^\top X_t$ ordered according to decreasing magnitude of eigenvalues
- $F_t = P_t \stackrel{\text{def}}{=} X_t \Gamma_t$ PCs of X_t
- factors f , introduce noise u_i , i.e.
 $y_j = b_{j1}f_1 + b_{j2}f_2 + \dots + b_{jk}f_k + u_j$ or $Y_t = F_t B_t^\top + U_t$
- then $\Sigma_t = \text{Var}(X_t) = \text{Var}(F_t B_t^\top) + \text{Var}(U_t) = B_t \Delta_t B_t^\top + \Omega_t$,
 $\Delta_t = \text{Var}(F_t)$ diagonal matrix of PC variances at t

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Dynamic Conditional Correlations Model

Assume: $r_t | \mathcal{F}_{t-1} \sim N(0, D_t R_t D_t)$, $\varepsilon_t = D_t^{-1} r_t$,

$$D_t^2 = \text{diag}(\omega_i) + \text{diag}(\alpha_i) \odot r_{t-1} r_{t-1}^\top + \text{diag}(\beta_i) \odot D_{t-1}^2,$$

$$Q_t = S \odot (\mathbf{1}\mathbf{1}^\top - A - B) + A \odot \{P_{t-1} \varepsilon_{t-1} \varepsilon_{t-1}^\top P_{t-1}\} + B \odot Q_{t-1},$$

$$R_t = \{\text{diag}(Q_t)\}^{-1} Q_t \{\text{diag}(Q_t)\}^{-1}$$

where $r_t \in \mathbb{R}^P$, $D_t = \text{diag}(\sigma_{it}) \in \mathbb{R}^{P \times P}$, $\varepsilon_t \in \mathbb{R}^P$ standardized returns with $\varepsilon_{it} \stackrel{\text{def}}{=} r_{it} \sigma_{it}^{-1}$, $\mathbf{1}$ vector of ones; $P_{t-1} \stackrel{\text{def}}{=} \{\text{diag}(Q_t)\}^{1/2}$, $\omega_i, \alpha_i, \beta_i, A, B$ coefficients, \odot Hadamard (elementwise) product

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The DCC Model - Continued

- correlation targeting: $S = (1/T) \sum_{t=1}^T \varepsilon_t \varepsilon_t^\top$
- $Q_0 = \varepsilon_0 \varepsilon_0^\top$ positive definite, each subsequent Q_t also positive definite
- consistent but inefficient estimates: the log-likelihood function

$$L(\theta, \phi) = -\frac{1}{2} \sum_{t=1}^T \left\{ n \log(2\pi) + 2 \log |D_t| + \log |R_t| + \varepsilon_t^\top R_t^{-1} \varepsilon_t \right\},$$

where θ parameters in D and ϕ additional correlation parameters in R [▶ Back](#)



The DCC Model - Continued

Re-write:

$$L(\theta, \phi) = L_V(\theta) + L_C(\theta, \phi),$$

with volatility part $L_V(\theta)$ and correlation part $L_C(\theta, \phi)$,

$$\begin{aligned} L_V(\theta) &= -\frac{1}{2} \sum_{t=1}^T \left\{ n \log(2\pi) + \log |D_t|^2 + r_t^\top D_t^{-2} r_t \right\} \\ &= -\frac{1}{2} \sum_{t=1}^T \sum_{i=1}^d \left\{ \log(2\pi) + \log(\sigma_{it}^2) + \frac{r_{it}^2}{\sigma_{it}^2} \right\}, \end{aligned}$$

$$L_C(\theta, \phi) = -\frac{1}{2} \sum_{t=1}^T \left\{ \log |R_t| + \varepsilon_t^\top R_t^{-1} \varepsilon_t - \varepsilon_t^\top \varepsilon_t \right\}. \quad \text{Back}$$



Cornish-Fisher VaR Optimization

Log returns $X_t \in \mathbb{R}^p$:

$$\begin{aligned} & \underset{w \in \mathbb{R}^d}{\text{minimize}} && W_t \{-q_\alpha(w_t) \cdot \sigma_p(w_t)\} \\ & \text{subject to} && w_t^\top \mu = \mu_p, \quad w_t^\top \mathbf{1} = 1, \quad w_{t,i} \geq 0 \end{aligned}$$

here $W_t \stackrel{\text{def}}{=} W_0 \cdot \prod_{j=1}^{t-1} w_{t-j}^\top (1 + X_{t-j})$, \tilde{w} , W_0 initial wealth,

$$\sigma_p^2(w) \stackrel{\text{def}}{=} w_t^\top \Sigma_t w_t,$$

$$q_\alpha(w) \stackrel{\text{def}}{=} z_\alpha + (z_\alpha^2 - 1) \frac{S_p(w)}{6} + (z_\alpha^3 - 3z_\alpha) \frac{K_p(w)}{24} - (2z_\alpha^3 - 5z_\alpha) \frac{S_p(w)^2}{36},$$

here $S_p(w)$ skewness, $K_p(w)$ kurtosis, z_α is $N(0, 1)$ α -quantile. If $S_p(w)$, $K_p(w)$ zero, then obtain Markowitz allocation

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Risk Parity (Equal risk contribution)

Let $\sigma(w) = \sqrt{w^\top \Sigma w}$. Euler decomposition:

$$\sigma(w) \stackrel{\text{def}}{=} \sum_{i=1}^n \sigma_i(w) = \sum_{i=1}^n w_i \frac{\sigma(w)}{\partial w_i}$$

where $\frac{\sigma(w)}{\partial w_i}$ is the marginal risk contribution and $\sigma_i(w) = w_i \frac{\sigma(w)}{\partial w_i}$ the risk contribution of i-th asset. The idea of ERC strategy is to find risk balanced portfolio, such that:

$$\sigma_i(w) = \sigma_j(w)$$

i.e. the risk contribution is the same for all assets of the portfolio

[▶ Return to "Benchmark Strategies"](#)



60/40 allocation strategy

60/40 portfolio allocation strategy implies the investing of 60% of the portfolio value in stocks (often via a broad index such as S&P500) and 40% in government or other high-quality bonds, with regular rebalancing to keep proportions steady.

▶ [Return to "Benchmark Strategies"](#)



Portfolio Skewness and Kurtosis

Skewness S_P and excess kurtosis K_P are given by moment expressions

$$S_P(w) = \frac{1}{\sigma_P^3(w)} (m_3 - 3m_2m_1 + 2m_1^3)$$

$$K_P(w) = \frac{1}{\sigma_P^4(w)} (m_4 - 4m_3m_1 + 6m_2m_1^2 + 3m_1^4) - 3$$

where portfolio non-central moments also depend on w :

$$m_1 = \mu_P(w) \stackrel{\text{def}}{=} w^\top \mu$$

$$m_2 = \sigma_P^2 + m_1^2$$

$$m_3 = \sum_{i=1}^d \sum_{j=1}^d \sum_{k=1}^d w_i w_j w_k S_{ijk}$$



Portfolio Skewness and Kurtosis - Continued

$$m_4 = \sum_{i=1}^d \sum_{j=1}^d \sum_{k=1}^d \sum_{l=1}^d w_i w_j w_k w_l K_{ijkl},$$

where $\sigma_p^2(w) = w^\top \Sigma w$ and $S_{ijk} = E(r_i \times r_j \times r_k)$, $K_{ijkl} = E(r_i \times r_j \times r_k \times r_l)$ can be computed via sample averages from returns data.

S_{ijk} , K_{ijkl} determine the d -dimensional portfolio co-skewness and co-kurtosis tensors

$$S \stackrel{\text{def}}{=} \{S_{ijk}\}_{i,j,k=1,\dots,d} \in \mathbb{R}^{d \times d \times d}$$
$$K \stackrel{\text{def}}{=} \{K_{ijkl}\}_{i,j,k,l=1,\dots,d} \in \mathbb{R}^{d \times d \times d \times d}.$$

▶ Back



Regularity Conditions for Adaptive Lasso QR

A1 Sampling and smoothness: $\forall x$ in the support of X_i , $\forall y \in \mathbb{R}$,
 $f_{Y_i|X_i}(y|x)$, $f \in \mathcal{C}^k(\mathbb{R})$, $|f_{Y_i|X_i}(y|x)| < \bar{f}$, $|f'_{Y_i|X_i}(y|x)| < \bar{f}'$; $\exists \underline{f}$,
 such that $f_{Y_i|X_i}(x^\top \beta_\tau | x) > \underline{f} > 0$

A2 Restricted identifiability and nonlinearity: let $\delta \in \mathbb{R}^p$,
 $T \subset \{0, 1, \dots, p\}$, δ_T such that $\delta_{Tj} = \delta_j$ if $j \in T$, $\delta_{Tj} = 0$ if
 $j \notin T$; $T = \{0, 1, \dots, s\}$, $\bar{T}(\delta, m) \subset \{0, 1, \dots, p\} \setminus T$, then
 $\exists m \geq 0$, $c \geq 0$ such that

$$\inf_{\delta \in A, \delta \neq 0} \frac{\delta^\top \mathbb{E}(X_i X_i^\top) \delta}{\|\delta_{T \cup \bar{T}(\delta, m)}\|^2} > 0, \quad \frac{3\underline{f}^{3/2}}{8\bar{f}'} \inf_{\delta \in A, \delta \neq 0} \frac{\mathbb{E}[|X_i^\top \delta|^2]^{3/2}}{\mathbb{E}[|X_i^\top \delta|^3]} > 0,$$

where $A \stackrel{\text{def}}{=} \{\delta \in \mathbb{R}^p : \|\delta_{T^c}\|_1 \leq c \|\delta_T\|_1, \|\delta_{T^c}\|_0 \leq n\}$



Regularity Conditions - Continued

A3 Growth rate of covariates:

$$\frac{q^3 \{\log(n \vee p)\}^{2+\eta}}{n} \rightarrow 0, \eta > 0$$

A4 Moments of covariates: Cramér condition

$$E[|x_{ij}|^k] \leq 0.5 C_m M^{k-2} k!$$

for some constants $C_m, M, \forall k \geq 2, j = 1, \dots, p$

A5 Well-separated regression coefficients: $\exists b_0 > 0$, such that
 $\forall j \leq q, |\hat{\beta}_j| > b_0$



Proof of the CF-CVaR Expansion 1

Define the Cornish-Fisher expansion:

$$q_{1-\alpha} \stackrel{\text{def}}{=} z_{1-\alpha} + (z_{1-\alpha}^2 - 1)s + (z_{1-\alpha}^3 - 3z_{1-\alpha})k - (2z_{1-\alpha}^3 - 5z_{1-\alpha})s^2,$$

where $s \stackrel{\text{def}}{=} S/6$, $k \stackrel{\text{def}}{=} K/24$, S and K are skewness and excess kurtosis, respectively; $z_{1-\alpha} \stackrel{\text{def}}{=} \Phi^{-1}(1 - \alpha)$.

Re-write:

$$q_{1-\alpha} = a_0 + a_1 z_{1-\alpha} + a_2 z_{1-\alpha}^2 + a_3 z_{1-\alpha}^3, \quad (19)$$

where $a_0 = -s$, $a_1 = 1 - 3k + 5s^2$, $a_2 = s$, $a_3 = k - 2s^2$



Proof of the CF-CVaR Expansion 2

Define the *conditional Value-at-Risk* (CVaR) or *expected shortfall* (ES):

$$\text{CVaR}_\alpha \stackrel{\text{def}}{=} \frac{1}{1-\alpha} \int_\alpha^1 \text{VaR}_q \, d q,$$

where $\text{VaR}_q \stackrel{\text{def}}{=} -\Phi^{-1}(\alpha)\sigma\sqrt{T}$

Observe:

$$\text{CVaR}_\alpha = -\frac{1}{1-\alpha} \int_\alpha^1 \Phi^{-1}(\alpha)\sigma\sqrt{T} \, d q \quad (20)$$

$$= -\frac{\sigma\sqrt{T}}{1-\alpha} \int_\alpha^1 \Phi^{-1}(\alpha) \, d q \quad (21)$$

$$= -\frac{\sigma\sqrt{T}}{1-\alpha} \int_{-\infty}^{-z_\alpha} u\varphi(u) \, d u, \quad (22)$$

where (22) follows from the change of variable: $u = z_q = \Phi^{-1}(q)$



Proof of the CF-CVaR Expansion 3

Substitute (19) into (22):

$$\begin{aligned}\text{CVaR}_\alpha &= -\frac{\sigma\sqrt{T}}{1-\alpha} \int_{-\infty}^{-z_\alpha} (a_0 + a_1z + a_2z^2 + a_3z^3) \varphi(z) dz \\ &= a_0A_0 + a_1A_1 + a_2A_2 + a_3A_3,\end{aligned}$$

$$A_0 = -\frac{\sigma\sqrt{T}}{1-\alpha} \int_{-\infty}^{-z_\alpha} \varphi(z) dz = -\sigma\sqrt{T},$$

$$A_1 = -\frac{\sigma\sqrt{T}}{1-\alpha} \int_{-\infty}^{-z_\alpha} z\varphi(z) dz = \frac{\sigma\sqrt{T}}{1-\alpha} \varphi(z_\alpha),$$

$$A_2 = -\frac{\sigma\sqrt{T}}{1-\alpha} \int_{-\infty}^{-z_\alpha} z^2\varphi(z) dz = -\sigma\sqrt{T} \left(\frac{\varphi(z_\alpha)z_\alpha}{1-\alpha} + 1 \right),$$

$$A_3 = -\frac{\sigma\sqrt{T}}{1-\alpha} \int_{-\infty}^{-z_\alpha} z^3\varphi(z) dz = \frac{\sigma\sqrt{T}}{1-\alpha} (z_\alpha^2 + 2)\varphi(z_\alpha).$$

Collecting terms and simplifying gives the desired result. □

[Return to "Conditional VaR Optimization"](#)



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